

C3 P, E

P1) Tenemos $(X, Z) = h(X, Y) = \left(X, \frac{X}{X+Y}\right)$

a) h es injectiva:

$$\text{Sean } X_1, Y_1, X_2, Y_2 \in \mathbb{R}: h(X_1, Y_1) = h(X_2, Y_2)$$

$$\Rightarrow \left(X_1, \frac{X_1}{X_1+Y_1}\right) = \left(X_2, \frac{X_2}{X_2+Y_2}\right)$$

$$\Rightarrow X_1 = X_2 \quad \wedge \quad \frac{X_1}{X_1+Y_1} = \frac{X_2}{X_2+Y_2}$$

$$\Rightarrow \frac{1}{X_1+Y_1} = \frac{1}{X_2+Y_2}$$

$$\Rightarrow X_1 + Y_1 = X_2 + Y_2 \Rightarrow Y_1 = Y_2 //$$

• Por TCV:

$$f_{X,Z}(x,z) = |\det(J_h(h^{-1}(x,z)))|^{-1} f_{X,Y}(h^{-1}(x,z)) = |\det(J_{h^{-1}}(x,z))| f_{X,Y}(h^{-1}(x,z))$$

$$* h(x,y) = \left(x, \frac{x}{x+y}\right) \Rightarrow h^{-1}(x,z) = \left(x, \frac{x}{z} - x\right)$$

$$* J_{h^{-1}}(x,z) = \begin{pmatrix} 1 & \frac{1}{z} - 1 \\ 0 & -\frac{x}{z^2} \end{pmatrix} \Rightarrow \det(J_{h^{-1}}(x,z)) = -\frac{x}{z^2}$$

* X es el valor de una exp, Z es división de valores $\geq 0 \Rightarrow X, Z \geq 0$

$$\begin{aligned} \therefore f_{X,Z}(x,z) &= \left| -\frac{x}{z^2} \right| f_X(x) f_Y\left(\frac{x}{z} - x\right) = \frac{x}{z^2} \cdot \lambda e^{-\lambda x} \cdot \lambda e^{-\lambda\left(\frac{x}{z} - x\right)} \mathbb{1}_{x>0} \mathbb{1}_{\frac{x}{z} - x > 0} \\ &= \frac{x}{z^2} \lambda e^{-\lambda x} \lambda e^{-\lambda \frac{x}{z}} e^{\lambda x} \mathbb{1}_{x>0} = \frac{x}{z^2} \lambda^2 e^{-\lambda \frac{x}{z}} \mathbb{1}_{x>0} \mathbb{1}_{0 < z < 1} \end{aligned}$$

* Usamos que X, Y indep

* Notemos $\frac{x}{z} - x > 0 \Rightarrow \frac{x}{z} > x$. Como $x > 0 \Rightarrow z > 0$

$$\text{Luego } \frac{1}{z} > 1 \Rightarrow z < 1 \quad \therefore 0 < z < 1$$

P1) b) Calculamos $f_Z(z)$:

$$f_Z(z) = \int_{-\infty}^{+\infty} f_{X,Z}(x,z) dx = \int_{-\infty}^{+\infty} \lambda^2 \frac{x}{z^2} e^{-\lambda \frac{x}{z}} \mathbb{1}_{x>0} \mathbb{1}_{0<z<1} dx$$

$$= \frac{\lambda^2}{z} \mathbb{1}_{0<z<1} \int_0^{+\infty} \frac{x}{z} e^{-\lambda \frac{x}{z}} dx$$

$$\text{Def } u = \frac{x}{z} \Rightarrow du = \frac{1}{z} dx$$

$$\therefore f_Z(z) = \frac{\lambda^2}{z} \mathbb{1}_{0<z<1} \int_0^{+\infty} u e^{-\lambda u} z du = \lambda^2 \mathbb{1}_{0<z<1} \int_0^{+\infty} u e^{-\lambda u} du$$

$$= \lambda \mathbb{1}_{0<z<1} \underbrace{\int_0^{+\infty} u \lambda e^{-\lambda u} du}_{\mathbb{E}(U), \text{ con } U \sim \text{Exp}(\lambda)} = \lambda \mathbb{1}_{0<z<1} \cdot \frac{1}{\lambda} = \mathbb{1}_{0<z<1}$$

$\mathbb{E}(U)$, con $U \sim \text{Exp}(\lambda)$

$$\therefore f_Z(z) = \mathbb{1}_{0<z<1} \Rightarrow Z \sim \text{Unif}(0,1) \quad \square$$

P2) a)
$$P(\forall i=1, \dots, n: U_i \in (s, t]) \stackrel{\text{indep}}{=} \prod_{i=1}^n P(U_i \in (s, t]) = \prod_{i=1}^n [F_{U_i}(t) - F_{U_i}(s)]$$

$$= \prod_{i=1}^n [t-s] = (t-s)^n$$

b) Notemos que el evento:

$\{S > s\} \Leftrightarrow \{\min(U_1, \dots, U_n) > s\} \Leftrightarrow \{\forall i=1, \dots, n: U_i > s\}$

$\{T \leq t\} \Leftrightarrow \{\max(U_1, \dots, U_n) \leq t\} \Leftrightarrow \{\forall i=1, \dots, n: U_i \leq t\}$

luego $P(S > s, T \leq t) = P(\forall i=1, \dots, n: U_i > s, U_i \leq t) = P(\forall i=1, \dots, n: U_i \in (s, t]) = (t-s)^n$

c) Dado $0 \leq s < t \leq 1$ fijos

Por prob totales: $P(T \leq t) = P(T \leq t, S > s) + P(T \leq t, S \leq s)$
 $\Rightarrow P(T \leq t, S \leq s) = P(T \leq t) - P(T \leq t, S > s)$

luego $P(T \leq t) = P(\forall i=1, \dots, n: U_i \leq t) = \prod_{i=1}^n P(U_i \leq t) = \prod_{i=1}^n F_{U_i}(t) = \prod_{i=1}^n t = t^n$

$\therefore P(T \leq t, S \leq s) = t^n - (t-s)^n$

• Por lo tanto $F_{S,T}(s, t) = t^n - (t-s)^n$

Por def $f_{S,T}(s, t) = \frac{\partial^2 F_{S,T}}{\partial s \partial t}(s, t) = \frac{\partial}{\partial t} \left(\frac{\partial F_{S,T}}{\partial s} \right)(s, t)$

$$= \frac{\partial}{\partial t} \left(n(t-s)^{n-1} \right) = n(n-1)(t-s)^{n-2}$$

$\therefore$ Para $0 \leq s < t \leq 1$

$f_{S,T}(s, t) = n(n-1)(t-s)^{n-2}$

P3) a) $E(e^{tx}) = \int_{-\infty}^{+\infty} e^{tx} f_X(x) dx = \int_{-\infty}^{+\infty} e^{tx} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2}} dx = \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \frac{(x-\mu)^2}{\sigma^2} + tx} dx$

$$= \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{(x-\mu)^2 + 2\sigma^2 tx}{2\sigma^2}} dx$$

• Trabajemos el término dentro de la exponencial:

$$\begin{aligned} 2\sigma^2 tx - (x-\mu)^2 &= 2\sigma^2 tx - x^2 + 2x\mu - \mu^2 = -x^2 + 2(\sigma^2 t + \mu)x - \mu^2 \\ &= -x^2 + 2(\sigma^2 t + \mu)x - \mu^2 + (\sigma^2 t + \mu)^2 - (\sigma^2 t + \mu)^2 \\ &= -[x^2 - 2(\sigma^2 t + \mu)x + (\sigma^2 t + \mu)^2] + (\sigma^2 t + \mu)^2 - \mu^2 \\ &= -[x - (\sigma^2 t + \mu)]^2 + (\sigma^2 t + \mu)^2 - \mu^2 \end{aligned}$$

• $E(e^{tx}) = \int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{[x - (\sigma^2 t + \mu)]^2 + (\sigma^2 t + \mu)^2 - \mu^2}{2\sigma^2}} dx$

$$= e^{\frac{(\sigma^2 t + \mu)^2 - \mu^2}{2\sigma^2}} \underbrace{\int_{-\infty}^{+\infty} \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2} \frac{(x - (\sigma^2 t + \mu))^2}{\sigma^2}} dx}_{=1, \text{ pues es igual a } \int_{-\infty}^{+\infty} f_Y(y) dy, \text{ con } Y \sim N(\sigma^2 t + \mu, \sigma^2)}$$

$$\begin{aligned} \Rightarrow E(e^{tx}) &= e^{\frac{(\sigma^2 t + \mu)^2 - \mu^2}{2\sigma^2}} = e^{\frac{\sigma^4 t^2 + 2\sigma^2 t\mu + \mu^2 - \mu^2}{2\sigma^2}} = e^{\frac{\sigma^4 t^2 + 2\sigma^2 t\mu}{2\sigma^2}} = e^{\frac{\sigma^2 t^2 + 2t\mu}{2}} \\ &= e^{t\left(\frac{\sigma^2}{2} + \mu\right)} \end{aligned}$$

P3) b) Usando la parte (a):

$$\bullet E(Y) = E(e^X) = e^{\frac{\sigma^2}{2} + \mu}$$

$$\bullet \text{Var}(Y) = E(Y^2) - [E(Y)]^2$$

Calculamos $E(Y^2)$:

$$E(Y^2) = E((e^X)^2) = E(e^{2X}) = e^{2(\frac{\sigma^2}{2} + \mu)}$$

$$\begin{aligned} \Rightarrow \text{Var}(Y) &= e^{2(\frac{\sigma^2}{2} + \mu)} - \left(e^{\frac{\sigma^2}{2} + \mu} \right)^2 = e^{2(\frac{\sigma^2}{2} + \mu)} - e^{2(\frac{\sigma^2}{2} + \mu)} = e^{2\sigma^2 + 2\mu} - e^{\sigma^2 + 2\mu} \\ &= e^{2\sigma^2} \cdot e^{2\mu} - e^{\sigma^2} \cdot e^{2\mu} = e^{2\mu} (e^{2\sigma^2} - e^{\sigma^2}) \end{aligned}$$