

PREGUNTA 1

$$1.1 \quad Y = \frac{X}{a} \Rightarrow F_Y(y) = P(Y \leq y) = P(X \leq ay) = F_X(ay) \Rightarrow$$

$$f_Y(y) = af_X(ay) = a\beta e^{-a\beta y} \Rightarrow Y \sim \exp(a\beta)$$

$$1.2 \quad L = \beta^n e^{-\beta \sum x_i} \Rightarrow \log(L) = n \log(\beta) - \beta \sum x_i$$

$$\frac{\partial \log(L)}{\partial \beta} = 0 \Rightarrow \frac{n}{\hat{\beta}} = \sum x_i \Rightarrow \hat{\beta} = \frac{1}{\bar{x}}.$$

$$1.3 \quad X_i \sim N(1, 4) \Rightarrow \bar{x} \sim N(1, \frac{4}{n}). \text{ Si } Z \sim N(0, 1) \Rightarrow P(|Z| \leq 1.96) = 0.95$$

$$\Rightarrow P(\bar{x} \in [1 - 1.96 \frac{2}{\sqrt{n}}, 1 + 1.96 \frac{2}{\sqrt{n}}]) = 0.95.$$

1.4 Si se repiten 100 muestras, se esperan que 95 de las muestras tengan su media muestral que cae en el intervalo.

PREGUNTA 2

$$2.1 \quad E(\hat{\mu}) = \sum_{i=1}^n a_i E(X_i) + \sum_{j=1}^m b_j E(Y_j) = \mu (\sum_{i=1}^n a_i + \sum_{j=1}^m b_j) = \mu \Rightarrow \boxed{\sum_{i=1}^n a_i + \sum_{j=1}^m b_j = 1}$$

$$2.2 \quad V(\hat{\mu}) = \sum_{i=1}^n a_i^2 \text{Var}(X_i) + \sum_{j=1}^m b_j^2 \text{Var}(Y_j) = \sigma^2 (\sum_{i=1}^n a_i^2 + 2 \sum_{j=1}^m b_j^2). \text{ Hay que minimizar}$$

$$\text{Var}(\hat{\mu}) \text{ sujeto a } \sum_{i=1}^n a_i + \sum_{j=1}^m b_j = 1 \Rightarrow Q = \text{Var}(\hat{\mu}) - \lambda (\sum_{i=1}^n a_i + \sum_{j=1}^m b_j - 1)$$

$$\left. \begin{aligned} \frac{\partial Q}{\partial a_i} = 0 &\Rightarrow 2\sigma^2 a_i - \lambda = 0 \Rightarrow a_i = \frac{\lambda}{2\sigma^2} \quad (a_i \text{ todos iguales}) \\ \frac{\partial Q}{\partial b_j} = 0 &\Rightarrow 4\sigma^2 b_j - \lambda = 0 \Rightarrow b_j = \frac{\lambda}{4\sigma^2} \quad (b_j \text{ todos iguales}) \end{aligned} \right\} \Rightarrow a_i = 2b_j$$

$$\text{De } \sum_{i=1}^n a_i + \sum_{j=1}^m b_j = 1 \text{ se deduce que } \boxed{a_i = \frac{2}{2n+m}, b_j = \frac{1}{2n+m}} \text{ y}$$

$$\boxed{\hat{\mu} = \left(\frac{1}{2n+m} \right) (2 \sum_i X_i + \sum_j Y_j)}$$

Es decir que es una media ponderada de todas las

observaciones, en donde los X_i tiene un peso doble del peso de los Y_j .

2.3 Calculemos la cota inferior de la desigualdad de Cramer-Rao. La función de verosimilitud de las $(n+m)$ valores muestrales es:

$$L = \left(\frac{1}{\sqrt{2\pi\sigma^2}} \right)^n \exp\left(-\frac{1}{2\sigma^2} \sum_i (X_i - \mu)^2\right) \left(\frac{1}{\sqrt{4\pi\sigma^2}} \right)^m \exp\left(-\frac{1}{4\sigma^2} \sum_j (Y_j - \mu)^2\right)$$

$$\log(L) = -\left(\frac{n}{2}\right) \log(2\pi\sigma^2) - \left(\frac{m}{2}\right) \log(4\pi\sigma^2) - \frac{1}{\sigma^2} \left\{ \frac{1}{2} \sum_i (X_i - \mu)^2 + \frac{1}{4} \sum_j (Y_j - \mu)^2 \right\}$$

$$\frac{\partial \log(L)}{\partial \mu} = \frac{1}{\sigma^2} \left\{ \sum_i (X_i - \mu) + \frac{1}{2} \sum_j (Y_j - \mu) \right\}$$

$$\frac{\partial^2 \log(L)}{\partial \mu^2} = -\frac{1}{\sigma^2} \left\{ n + \frac{m}{2} \right\} = -\frac{1}{\sigma^2} \left\{ \frac{2n+m}{2} \right\} \Rightarrow \boxed{Var(\hat{\mu}) \geq \frac{2\sigma^2}{2n+m}}$$

Por otro lado:

$$Var(\hat{\mu}) = \sigma^2 \left(\sum_i a_i^2 + 2 \sum_j b_j^2 \right) = \frac{2\sigma^2}{2n+m} \Rightarrow \hat{\mu} \text{ es un estimador insesgado de mínima}$$

varianza para μ .

2.4 Se tiene $\mu^* = \sum_{i=1}^n a_i X_i + \sum_{j=1}^m b_j Y_j$ y $E\{(\mu^* - \mu)^2\} = Var(\mu^*) + \{E(\mu^* - \mu)\}^2$.

$$Var(\mu^*) = \sigma^2 \left(\sum_i a_i^2 + 2 \sum_j b_j^2 \right) \text{ y } E(\mu^* - \mu) = \mu \left(\sum_i a_i + \sum_j b_j - 1 \right)$$

$$E\{(\mu^* - \mu)^2\} = \sigma^2 \left(\sum_i a_i^2 + 2 \sum_j b_j^2 \right) + \mu^2 \left(\sum_i a_i + \sum_j b_j - 1 \right)^2$$

Minimizando:

$$\frac{\partial E\{(\mu^* - \mu)^2\}}{\partial a_i} = 2\sigma^2 a_i + 2\mu^2 \left(\sum_i a_i + \sum_j b_j - 1 \right) = 0 \Rightarrow a_i = \frac{\mu^2}{\sigma^2} \left(1 - \sum_i a_i - \sum_j b_j \right)$$

$$\frac{\partial E\{(\mu^* - \mu)^2\}}{\partial b_j} = 2\sigma^2 b_j + 2\mu^2 \left(\sum_i a_i + \sum_j b_j - 1 \right) = 0 \Rightarrow b_j = \frac{\mu^2}{2\sigma^2} \left(1 - \sum_i a_i - \sum_j b_j \right)$$

Luego todos los b_j son iguales y todos los a_i son iguales a $2b_j$. Sea $b_j = b$ y $a_i = 2b$.

Entonces $\sum_i a_i = \frac{n\mu^2}{\sigma^2} (1 - \sum_i a_i - \sum_j b_j)$ o sea $2nb = \frac{n\mu^2}{\sigma^2} (1 - 2nb - mb)$

Finalmente: $b = \frac{\mu^2}{2\sigma^2 + (2n+m)\mu^2}$ y $\mu^* = \sum_{i=1}^n a_i X_i + \sum_{j=1}^m b_j Y_j = 2b \sum X_i + b \sum Y_j$

$$\mu^* = b(2 \sum X_i + \sum Y_j) = \frac{\mu^2}{2\sigma^2 + (2n+m)\mu^2} (2 \sum X_i + \sum Y_j)$$

2.5 De aquí se deduce la varianza de μ^* :

$$Var(\mu^*) = \left(\frac{\mu^2}{2\sigma^2 + (2n+m)\mu^2} \right)^2 (4n\sigma^2 + m\sigma^2) = \frac{2(2n+m)\sigma^2}{(2\sigma^2 / \mu^2 + (2n+m))^2}$$

Como $Var(\hat{\mu}) = \frac{2\sigma^2}{2n+m} = \frac{2(2n+m)\sigma^2}{(2n+m)^2}$, se deduce que $\boxed{Var(\hat{\mu}) > Var(\mu^*)}$

PREGUNTA 3

$$3.1 \quad \xi(\theta | X) = \begin{cases} \frac{\theta + 2(1-\theta)x}{2g(x)} & \text{si } 0 \leq \theta \leq 2 \\ 0 & \text{si no} \end{cases} \quad \text{con } g(x) = \int_0^2 \frac{1}{2}(\theta + 2(1-\theta)x) d\theta = 1$$

$$\text{Luego } \xi(\theta | X) = \begin{cases} \frac{\theta + 2(1-\theta)x}{2} & \text{si } 0 \leq \theta \leq 2 \\ 0 & \text{si no} \end{cases}$$

$$3.2 \text{ Bajo perdida cuadrática, } \theta_B = E(\theta | x) = \int_0^2 \frac{\theta}{2}(\theta + 2(1-\theta)x) d\theta = \frac{4}{3} - \frac{2}{3}x$$

$$3.3 \quad E(\theta_B | x) = \frac{4}{3} - \frac{2}{3}E(x)$$

$$E(x) = \int_0^1 x(\theta + 2(1-\theta)x) dx = \theta \frac{x^2}{2} + 2(1-\theta) \frac{x^3}{3} \Big|_0^1 = \left(\frac{\theta}{2} + \frac{2(1-\theta)}{3} \right) = \frac{2}{3} - \frac{\theta}{6}$$

$$\text{Luego } E(\theta_B | x) = \frac{4}{3} - \frac{2}{3}E(x) = \frac{4}{3} - \frac{2}{3} \left(\frac{2}{3} - \frac{\theta}{6} \right) = \frac{\theta + 8}{9} \Rightarrow \text{el estimador es sesgado.}$$

$$3.4 \quad \text{Si } x=1/2,$$

$$P(0 \leq \theta \leq k | x = 1/2) = \int_0^k \xi(\theta | x = 1/2) d\theta = \int_0^k \frac{1}{2}(\theta(1-\theta)) d\theta = \int_0^k \frac{1}{2} d\theta = \frac{k}{2}$$

$$P(0 \leq \theta \leq k | x = 1/2) = 1/2 \Rightarrow \boxed{k=1}$$